

Efficiency of crosses selection for yield in red rice through diallel analysis

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ABSTRACT

Diallel analysis is a methodology used in the genetic improvement of plants. One of its main purposes is to provide subsidies for early identification of promising crosses for the development of high yielding inbred lines. Several articles have been published in diallel analysis, and many recommendations were made either for the selection of certain crosses of lines for hybrids or for the development of competitive inbred lines. However, articles about the validation of information predicted by these analyses and the results in the field are scarce or nonexistent. The objective of this article was to validate diallel analysis using model IV of Gardner and Eberhart published in 1966, a methodology recommended to identify promising crosses among developed inbred lines in red rice. We compared the recommendations of the diallel analysis crosses with the process end the result of inbreeding generations. Recommended crosses by the diallel analysis were different from those obtained after generations of self-fertilization and selection. However, two of the four selected parents were matched. Considerations were made about the accuracy of diallel analysis and the need to validate in practice methodologies that theoretically are of great value.

Keywords: Reliability, plant breeding, prediction, diallelic cross, parent selection, combination ability.

INTRODUCTION

Diallelic cross analysis is a widely used methodology for the genetic improvement of plants (Chukwu, Okporie, Onyishi, Ekwu, & Nwogbaga, 2016; Fasahat, Rajabi, Rad, & Derera, 2016). In the study of the gene effects of the characters, more specifically in allogamous plants, it provides subsidies in the selection of parents to obtain commercial hybrids. In autogamous plants, it is also useful for the early identification of promising crosses aiming at the development of new inbred lines with more possibilities to become cultivars (Vencovsky & Barriga, 1992; Cruz et al., 2012).

Diallel analysis methodology consists of the study of a set of hybrids that may include the respective parents, segregant generations as F_2' , F_3' , etc., backcrosses and others. Common diallel analysis procedure are Hayman (1954a), Hayman (1954b), Griffing (1956) and Gardner and Eberhart (1966). Evaluation of crosses allows the estimation of the value per se of the parents and their combinatorial abilities, enabling the identification of the best populations derived from them regarding the genetic diversity (Hallauer, Carena, & Miranda-Filho, 2010).

Several articles have been published using diallel analysis, and many recommendations were made for the selection of certain crosses of inbred lines in order to obtain commercial hybrids or to develop competitive pure lines (Borghini, & Perenzin, 1994; Mutimaamba et al., 2017). However, articles about the confrontation of predicted information from these analyzes and from those observed after the generations of self-fertilization are scarce or nonexistent. Nevertheless, there are still important questions and some restrictions on the diallel use as a source of information that may contribute to breeding programs (Baker, 1978). The objective of this paper was to validate the recommendations of promising crosses of a diallel analysis using the methodology of model IV of Gardner and Eberhart (1966), performed in red rice (Figure 1) by Pereira, Morais and Bresghello (2008).



Figure 1. Illustrative image of the red rice grains used in the experiment.
Photo: By Maria Eugenia Ribeiro, Embrapa Meio-Norte.

MATERIALS AND METHODS

Ten most commonly grown cultivars of red rice from the Northeast region of Brazil (named parents), three from the state of Alagoas (AL01, AL03, and AL04) and seven from the state of Paraiba (PB04, PB05, PB09, PB10, PB11, PB12 and PB13), were crossed in a complete 10 x 10 diallel design, without reciprocals, for the implementation of a breeding program. Parents and their 45 F_1 were evaluated in a field experiment in the second half of 2005, while parents and their 45 F_3 in the second half of 2006 regarding their agronomic characteristics. All two field experiments were in a randomized complete block design with three replicates in Teresina, Piauí state, Brazil (5°09' S; 42°48' W, considered an tropical wet and dry climate with an altitude of 72 m and the following average characteristics in the second semester: temperature 28 °C, relative humidity 63% and total rainfall 201 mm). For the diallel analysis, the ten parents and an additional cultivar (PB 01) were used as common treatments in the experimental evaluations involving the F_1 and F_3 generations.

F_1 generation experiment was grown under controlled flood irrigation with water flood control, in one-line plots with 12 transplanted seedlings, spaced 0.25 m between plants and 0.30 m between rows, whose sample area was the ten central plants of the line (0.75 m²). F_3 experiment was carried out under irrigated conditions using direct seeding with plots of three 5m rows spaced 0.30 m at a density of 20 seeds per linear meter. This special condition in the development of the F_1 generation was necessary to avoid plant losses due to the difficulty in obtaining F_1 crosses in autogamous plants. The favorable condition was corrected in the statistical analysis by adjusting the use of common treatments, as pointed out previously. The inbreeding generations process (F_2 , F_3 , etc.) were grown twice a year, in the rainy season and in the dry season with flood irrigation and a plot of 1.5 m² per treatment was used.

Diallel analysis of grain yield using model IV of Gardner and Eberhart (1966) methodology resulted in recommendations on promising crosses in the generation of high-yielding inbred lines described in Pereira et al. (2008). However, prior to this analysis, a soft selection of crosses for advancement to the F_4 generation was initiated in the F_3 (in 2006). Subsequently, the process of advancing generations through the pedigree method with a mild selection for smaller plants was followed, ignoring the recommendations of the diallel analysis, whose results were ready only two years after the F_3 generation. Finally, in the F_8 generation, two lines were identified to assess their value for cultivation and use.

RESULTS AND DISCUSSION

Table 1 shows the analysis of the variance of the diallel analysis partitioned into the main sources of variation. Diallel analysis of ten parental red rice carried out by Pereira et al. (2008) concluded that the greatest effects of general combining ability were on AL03, PB05 and PB13 parents and the promising crosses for the generation of new high-yielding red rice lines would be PB05 x PB10 and AL03 x PB13. However, a different result was found at the end of the inbreeding generation, since the selected crosses that originated a pure line of value for cultivation and use assays were PB04 x PB10 and PB05 x PB09, as presented in Table 2.

Table 1. Analysis of the variance of the diallel analysis of grain yield in g (m²)⁻¹ partitioned into the main sources of variation (Pereira et al., 2008).

Source of variation	Degrees of freedom	Sum of squares	Mean squares	F
Treatments	99	10.696.314	108.044	3.27**
Cultivars	9	656.860	72.984	2.21**
Heterosis in F₁	45	3.134.716	69.660	2.11**
Average heterosis in F₁	1	344.509	344.509	10.44**
Varietal heterosis in F₁	9	1.597.049	177.450	5.38**
Specific heterosis in F₁	35	2.677.344	76.496	2.32**
Heterosis in F₃	45	2.384.444	52.988	1.61**
Average heterosis in F₃	1	20.430	20.430	0.62 ^{ns}
Varietal heterosis in F₃	9	614.355	68.262	2.07**
Specific heterosis in F₃	35	1.709.850	48.853	1.48*
Residue	308	6.864.699	33.003	-

*Significant at the .05 probability level. **Significant at the .01 probability level. ^{ns}not significant.

The choice of the promising crosses for diallel analysis followed the recommendations in the literature (Gardner & Eberhart, 1966; Vencovsky & Barriga, 1992; Cruz, Regazzi, & Carneiro, 2012), in which the value of *g* (with its components (½)*v* and *h*) and parental productivity are presented in Table 3. The *v* effect refers to the value per se, and *h* effect is related to the diversity of the parents and the effect of gene dominance.

Genetic analysis of crosses considered the most promising by the diallel analysis

Selection of the AL03 x PB13 cross was made according to the theoretical recommendations for diallel analysis. These parents have the greatest positive deviations of the general combining ability (*g*), and were statistically different from zero, as well as the most productive (Table 3). In turn, in the selection of the PB05 x PB10 cross, it was verified that PB05 has the third largest value of *g*. It was statistically different from zero, and also the third place in productivity. However, PB10 was not an interesting parent, since its *g* ranked seventh and sixth in productivity. In addition, because of its *g* was not significant, it can be inferred that no evidence was found that its effect is different from zero (Christensen, 1996; Gomes, 2009). However, the choice of PB10 as a pair of PB05 was due to its excellent performance in specific heterosis with PB5, having reached the second largest positive and important deviations in F₁ and F₃ (Table 3).

Further clarification on the choice of promising crosses was needed by taking advantage of the partitioning of the general combining ability of the parents into their components (½) *v* (cultivar effect) and varietal *h* (heterosis effect: difference between the mean of the varietal and the general heterosis) (Table 3).

Table 2. Evolution of selections in segregating self-fertilizing generations.

F ₃	F ₄	F ₅	F ₆	F ₇	F ₈
	AL01 x (AL03, AL04, PB04, PB05, PB10)	AL01 x PB04	AL01 x PB10, 4 plants	-	-
	AL03 x (AL04, PB04, PB05, PB09, PB10)	AL01 x PB10	-	-	-
	AL04 x (PB04, PB05, PB09, PB10, PB11)	AL04 x PB4	AL04 x PB04, 5 plants	-	-
		AL04 x PB5	AL04 x PB5, 7 plants		-
		AL04 x PB09	-	-	-
		AL04 x PB10	AL04 x PB10, 4 plants	-	-
	PB04 x (PB10)	PB04 x PB10	PB04 x PB10, 3 plants	PB04 x PB10, 5 plants	1 inbred line
	PB05 x PB09	PB05 x PB09	PB05 x PB09, 3 plants	PB05 x PB09, 5 plants	1 inbred line
	PB12 x PB04	PB12 x PB04	PB12 x PB04, 3 plants	-	-
	PB13 x (AL01, AL03, PB05)	PB13 x AL01	PB13 x AL01, 2 plants	-	
45 crosses	21 crosses	10 crosses.	8 crosses (31plants)	2 crosses (10plants)	2 crosses (2 inbred lines)

Table 3. General combining ability gene effects (*g*) and yield partitioning in g (m⁻²)⁻¹ in parents of promising crosses based on Pereira et al. (2008).

Crosses	Parents	<i>g</i>	$\frac{1}{2} c$	<i>h</i> _{F1}	<i>h</i> _{F3}	<i>h</i>	$\frac{1}{2} c \%$	<i>h</i> %
AL03 x PB13	AL03	71.8**	85.2*	104.1*	-131.0*	-13.4	86	14
	PB13	66.7*	82.7*	38.3 ^{ns}	-70.5 ^{ns}	-16.1	84	16
PB05 x PB10	PB5	64.7*	39.3 ^{ns}	33.6 ^{ns}	17.2 ^{ns}	25.4	61	39
	PB10	19.8 ^{ns}	-0.9 ^{ns}	-22.8 ^{ns}	64.2 ^{ns}	20.7	4	96

c cultivar effect, *h*_{F1} heterosis effect in F₁, *h*_{F3} heterosis effect in F₃, *h* heterosis effect calculated as $\frac{1}{2} (h_{F1} + h_{F3})$, $\frac{1}{2} c \%$ $\frac{1}{2} c$ as a percentage of the *g* effect, *h*% *h* as a percentage of the *g* effect. ^{ns}, * and **; nonsignificant and significant at 5 and 1% probability by the t test, respectively.

Cross AL03 x PB13

In agreement with the significance of *g* of the parents of this cross, their estimates of ($\frac{1}{2}$) *v* are also positive, of great magnitude and statistically different from zero. In addition, varietal heterosis effect of AL03 is greatly and of the greatest magnitude in both F₁ and F₃, although it assumes opposite directions, on the positive side on F₁ and negative on F₃. Thus, the contribution of *h* to *g* in this parent is only 14% and remains 86% for $\frac{1}{2} v$, which is theoretically very interesting due to *v* related to the additive effects of the genes, reinforcing the choice of this parent. In the case of PB13 parent, the $\frac{1}{2} v$ effect is positive, of great magnitude and very similar to the first one, contributing with 84% of the value of *g*, whereas deviations of

varietal heterosis effect were not detected as significant in the precision conditions of the experiment. Thus, it is emphasized that at AL03 x PB13 cross, the specific heterosis in F_1 and F_3 were of small magnitude and not statically different from zero.

Cross PB05 x PB10

Analyzing the PB5 x PB10 cross, it was verified that in PB05 parent, despite statistical significance for g due to those infrequent situations, it was not possible to detect that its components are different from zero. Moreover, PB10, on the other hand, did not have a noticeably g effect and also its components followed the same trend. Considering these facts, it is estimated that the PB5 x PB10 crossing was due to the merits of PB05 regarding grain yield (third place), magnitude and significance of its g (also third place).

Genetic analysis of selection crosses with better performance on selection of promising lines in practice.

The crosses that, at the end of the generations by successive self-fertilization, generated inbred lines for Value for Cultivation and Use test were selected from the PB04 x PB10 and AL04 x PB05 crosses whose genetic analyzes are shown in Table 4.

Table 4. General combining ability (g) gene effects and partitioning of yield in g (m^2)⁻¹ of parents of crosses identified after generation of inbreeding and selection of promising inbred lines.

Crosses	Parents	g	$\frac{1}{2} c$	h_{F_1}	h_{F_3}	h	$\frac{1}{2} c \%$	$h\%$
PB04 x PB10	PB04	38.8 ^{ns}	-39.2 ^{ns}	56.0 ^{ns}	100.1*	78.0 ^{ns}	33	67
	PB10	19.8 ^{ns}	-0.9 ^{ns}	-22.8 ^{ns}	64.2 ^{ns}	20.7	4	96
AL04 x PB05	AL04	-136.1**	15.5 ^{ns}	-239.2**	-63.9 ^{ns}	-151.5	9	91
	PB5	64.7*	39.3 ^{ns}	33.6 ^{ns}	17.2 ^{ns}	25.4	61	39

c cultivar effect, h_{F_1} heterosis effect in F_1 , h_{F_3} heterosis effect in F_3 , h heterosis effect calculated as $\frac{1}{2} (h_{F_1} + h_{F_3})$, $\frac{1}{2} c \%$ as a percentage of the g effect, $h\%$ h as a percentage of the g effect. ^{ns}, * and **, nonsignificant and significant at 5 and 1% probability by the t test, respectively.

Cross PB04 x PB10

Diallel analysis showed that the g effects of parents of this cross were of small magnitude, not statistically different from zero, although the estimates were positive (38.8 g (m^2)⁻¹ and 19.8 g (m^2)⁻¹ respectively and the maximum g effect obtained was 71.8 g (m^2)⁻¹. In the partitioning of the g s effects, it was also found that the $\frac{1}{2} v$ effects also were not statistical significant. However, heterosis of F_3 (h_{F_3}) of PB04 was the largest and unique positive and statistically different from zero of the heterosis effects in this generation, and the percentage participation of h in g was 67% in this parent. This seems to indicate that PB04 is not attractive to be selected. This is due to a low g , grain yield basically dependent on varietal heterosis demonstrating its low accumulation of additive gene effects.

Cross AL04 x PB05

AL04 occupies the fourth position in yield productivity. It has a statistically significant g , however negative, and is theoretically not interesting due to the fact that this parent has a high magnitude of additive gene effects contrary to productivity. The partitioning of g indicates that heterosis in F_1 (h_{F_1}) has the greatest negative, highly marked deviation. When added to the effects of h_{F_3} , h contributes 91% to the g value, remaining only 9% for the additive gene effects (Table 3). Analyzing h in more details, in this case, we can see that its deviation h_{F_1} is highly noticeably, of greater negative magnitude among all of them. Its h_{F_3} , also negative. Based on this information, it can be said that, besides the fourth position in productivity, this parent has no other genetic characteristics revealed by the diallel analysis that can make it attractive to be selected.

The PB05, one of the parents selected at the crosses considered as promising by the diallel analysis, has been described as unattractive, according to the recommendations for the parent selection. But PB05 has some merits: it is a productive parent (third in the classification) with a good g effect, the third largest positive and noticeably g . However, it failed to demonstrate statistical significance at $\frac{1}{2} v$ and h components.

In the absence of agreement between the recommendations for the diallel analysis and the results of promising crosses selection after the generation of inbreeding, the selection of inbred lines valuable for cultivation and use tests, performed in practice, can be attributed to:

Selection of families performed during generations of the inbreeding process. In the F_3 generation, 47% of the crosses were chosen, excluding explicitly defective individuals. In the selection of F_4 , 48% of the remaining were chosen. In F_5 , 80% of the remnants were separated, being more common than that in the latter generation, where, in fact, the selection of plants was carried out. It is worth emphasizing that the effects of dominance can mask the vigor of the plants to be selected. This occurs especially in the first generations of self-fertilization, enabling the election of families with attributes to be diminished in subsequent inbreeding generations.

Another important drawback of early selection is the elimination of crosses in the first selections. In the present research, the promising crossover identified by the diallel analysis, PB05 x PB10 was eliminated from the inbreeding process in F_3 , and the other AL03 x PB13 was excluded in F_4 . However, it is important to emphasize that the attributes that guided early selection, such as excessive height of plants, very long cycle and type of grain, have high heritability known as unfavorable to the crop or preferred by the traditional consumer. Another aspect that justifies this selection is the fact that red rice breeding in Brazil is recent, and purification actions towards known harmful characteristics are necessary. Therefore, this can be justified by the soft selection practiced in the generations F_2 to F_4 .

Accuracy of the estimates that supported the recommendations of promising crosses is not optimal. These originated from an experiment with F_1 and parents in one year and one location, with parents and F_3 in another year, yet in the same location. More accurate diallel experiments involve several sites and two years, which, in practice, is difficult to perform due to difficulties in obtaining seeds of crosses in autogamous plants. With repeated experiments in places and years, it is possible to isolate the sources of genotype interaction variation within environments and years from the residue, as well as to improve the accuracy of estimates of the general and the specific combining ability of the crosses. Also, averages and deviations (g , $\frac{1}{2} v$, h) derive from a greater number of observations. In the experiment under consideration, the coefficient of variation was 22%, considered relatively high, and thus, its accuracy was considered low. Therefore, the relevance of the analyzed effects needs to be better evaluated by analyzing the relative dispersion of the magnitude of the estimated effects according to Cox (1982) (Tables 5 and 6). In addition, according to Christensen (1996), coefficient intervals are more valuable than simply reporting the estimates, since it provides an idea of the amount of error associated with the estimates.

Data in Table 5 show that the general combining abilities of parents of crosses considered to be promising by diallel analysis have moderately dispersions, around + -40% of their estimates (excluding PB10). Thus, the g of 'AL03' has a range of 45 $g (m^2)^{-1}$ to 100 $g (m^2)^{-1}$. The $\frac{1}{2} v$ components of AL03 and PB13 have dispersions similar to those of their g estimates, always presenting positive values, reinforcing the idea that the choice of this cross is adequate.

At the second cross selected in this group, PB05 x PB10, PB05 reveals dispersion for g also similarly in percentage to the parents of the first selected cross. However, when observing the partitioning of g , new information arises. The estimate of $\frac{1}{2} v$ has a very broad dispersion of + -98% of their estimate, ranging from uncomfortable 0.7 $g (m^2)^{-1}$ at the lower limit to 78 $g (m^2)^{-1}$ at the upper limit. The 'PB10', second parent of this cross, far exceeded the magnitude of its dispersion when comparing it with the first three parents analyzed. The dispersion of its g estimate is $\pm 136\%$ of their estimate, ranging from -7 $g (m^2)^{-1}$ to 47 $g (m^2)^{-1}$ and may assume negative values. In turn, the dispersion of its $\frac{1}{2} v$ estimate reaches exceptionally high amount, +3928% of their estimate, ranging from -63 $g (m^2)^{-1}$ to 79 $g (m^2)^{-1}$, thus can assume many negative values.

Next considerations on the dispersions of estimates of the general combining ability (and their components) of the parents of crosses effectively selected after self-fertilization generations are made (Table 6).

Cross AL04 x PB05

The CGA estimate of AL04 (-136.1 **) was highly significant and negative and with the lowest level of dispersion of all estimates in Tables 4 and 5 (Standard deviation of g percent $g = 21\%$). This would indicate the possession of genes that contribute very negatively to productivity which is not at all attractive in parent selection. Even worse, the dispersion of this CGA (g) only covers negative values ranging from $-165 \text{ g (m}^2\text{)}^{-1}$ to $-107 \text{ g (m}^2\text{)}^{-1}$. In the GCA partition, it is found that the estimate of $\frac{1}{2} v$ has a huge dispersion ($\pm 228\%$) and that can be influenced by the absence of statistical significance of g . We then used the effect of heterosis (h) to try to explain the high statistical significance of negative CGA of this cultivar. Its varietal heterosis in F_1 was highly negative ($-239.2 \text{ g (m}^2\text{)}^{-1}$), while its heterosis in F_3 was not notable (Pereira et al., 2008). Thus, the component h would explain the negative statistical significance of CGA. However, the low precision level of the F_3 heterosis effect may have had a negative influence on the detection of significance.

The parent, PB05, has been previously analyzed as being one of the parents of one of the crosses recommended by diallel analysis, whenever crossed with PB10. However, the presence of this parent, which has one of the best CGA, would give some theoretical support to the fact that the AL05 x Pb05 cross had originated the selection of an inbred line value for cultivation and use test in practice. Also, it suggests that the crosses selected by the diallel analysis under analysis, are not completely unfocused from the performance of crosses in the production of more productive lineages in the process of advancement of generations of inbreeding.

Cross PB04 x PB10

PB04 and PB10 are parents with relatively few merits. In productivity, they are below the group average. The diallel analysis showed that their CGAs were of small magnitude, not statistically different from zero, although positive $38.8 \text{ g (m}^2\text{)}^{-1}$ and $19.8 \text{ g (m}^2\text{)}^{-1}$ respectively, uncompetitive for a maximum obtained of $71.8 \text{ g (m}^2\text{)}^{-1}$. In the partitioning of CGA, it was also found that the v effects were not significant either. However, heterosis of F_3 (h_{F3}) from 'PB04' was the largest and single, positive and different from zero among evaluated genotypes and percentage participation of h in CGA was 67% in this genotype (Table 3). This seems to indicate that PB04 has a CGA low in magnitude and basically dependent on varietal heterosis. Such would demonstrate its low accumulation of additive effects. However, it is good to consider that the information of the dispersion of the g and v effects serve to indicate that the low level of precision of these estimates explains, in part, such absence of significance (Table 6).

Failure of assumptions of diallel analysis

Genetic interpretation of the results of diallel analysis is valid only if some assumptions about the parents are true. These assumptions were discussed by several authors (Kempthorne, 1956; Baker, 1978; Christie & Shattuck, 1992, among others). There are critical issues on diallelic experiments that meet some of these requirements. The following aspects are pointed out:

Establish that the genes are distributed independently is unsustainable (Kempthorne, 1956; Nassar, 1965). This assumption implies that the presence or absence of an allele at a particular locus is statistically independent of the presence or absence of an allele at any other locus. Failure of this assumption may result from genetic linkage in the population from which the parents were chosen. In turn, Baker (1978) concluded that assuming that genes are distributed independently in the parent of a diallel is unrealistic. However, he also said that there are contradictory reports of the effects of failures on estimates of both general and specific combining ability.

Absence of epistasis: According to Baker (1978) assumption of no epistasis can often be incorrect. Epistasis affects mean square estimates of general and specific combination ability, variances, and effects in an unpredictable manner. However, Nkalubo, Melis, Derera, Laing and Opio (2009) working with resistance to anthracnose in beans and not in yield in red rice, as in this case, were not conclusive about the action of epistasis in this character.

No genotype x environment interaction: According to Fasahat et al. (2016), this assumption is often not calculated to know the role of the genotype x environment interaction in combining ability. This can only be verified by evaluating diallel crosses in at least three environments, a requirement that was not observed in the present experiment. This leads to another practical aspect of using diallels as tools in selecting promising crosses in breeding programs, due to the high control cost of this assumption.

Table 5. General combining ability, the standard deviation of g in percent of g and dispersion of general combining ability (g) and $\frac{1}{2}$ cultivar effects ($\frac{1}{2} v$) of parents from promising crosses of a red rice diallel analysis of grain yield in $g (m^2)^{-1}$ based on Pereira et al., 2008.

Crosses	Parents	General combining ability (g)	Standard deviation of g in percent of g (sg%)	Dispersion (g -sg - g +sg)	$\frac{1}{2} c$	Standard deviation of $\frac{1}{2} c$ in percent of $\frac{1}{2} c$ ($s \frac{1}{2} c\%$) [§]	Dispersion ($\frac{1}{2} c$ - $s\frac{1}{2} c$ - $\frac{1}{2} c$ + $s\frac{1}{2} c$)
AL03 x	AL03	71.8**	38	45 – 100	85.2*	41	50 – 121
PB13	PB13	66.6*	40	40 – 93	82.7*	43	47 – 118
PB05 x	PB5	64.7*	40	39 – 91	39.3ns	98	0.7 – 78
PB10	PB10	19.8 ^{ns}	136	-7 – 47	-0.9ns	3928	-36 – 34

^{ns}, * and **; nonsignificant and significant at 5 and 1% probability by the t test.

[§] $s \frac{1}{2} c = [(\frac{1}{2})^2 s^2 \frac{1}{2} c]^{\frac{1}{2}}$.

Table 6. General combining ability, the standard deviation of g in percent of g and dispersion of general combining ability (g) and $\frac{1}{2}$ cultivar effects ($\frac{1}{2} v$) of parents of red rice crosses identified after generations of inbreeding and selection of inbred lines of value for cultivation and use test.

Crosses	Parents	General combining ability (g)	Standard deviation of g in percent of g (sg%)	Dispersion (g -sg - g +sg)	$\frac{1}{2} c$	Standard deviation of $\frac{1}{2} c$ in percent of $\frac{1}{2} c$ ($s \frac{1}{2} c\%$) [§]	Dispersion ($\frac{1}{2} c$ - $s\frac{1}{2} c$ - $\frac{1}{2} c$ + $s\frac{1}{2} c$)
AL04 x	AL04	136.1**	21	-165 – -	15.5*	228	-20 – 51
PB05				107			
	PB05	64.7*	40	39 – 91	39.3*	98	0.65 – 78
PB04 x	PB4	38.8 ^{ns}	65	13 – 64	-	90	-75 – 3.9
PB10					39.2 ^{ns}		
	PB10	19.8 ^{ns}	136	-7 – 47	-0.9 ^{ns}	3928	-36 – 34

^{ns}, * and **; nonsignificant and significant at 5 and 1% probability by the t test.

[§] $s \frac{1}{2} c = [(\frac{1}{2})^2 s^2 \frac{1}{2} c]^{\frac{1}{2}}$.

Restrictions on assumptions must be coupled with limitations on the accuracy of data that is taken in one place, at best, in a few of them and within one or two years. The reason for this is the reduction of costs, time, as well as the possibility of making the use of such predictions attractive. One could not spend much time, money, and effort to obtain data for predictions, if these items were equivalent to the use of generations' advancement methods without using such predictions.

The comment on the usefulness of the diallels is not new, since 1958 (Gilbert, 1958) had already concluded: "The diallel cross offers a means of rationalizing some aspects of plant breeding while keeping the amount of work down to a manageable level. Its utility to the breeder can be exaggerated. The polygene analysis of a diallel cross suffers from several theoretical defects, but in any case its results do not appear to be directly relevant to practical breeding work." Gilbert (1958) ends by indicating that the results

of a diallel analysis cannot be given as definitive, according to him: “No statistics can replace the breeder's intimate knowledge of his crop but they can help.”

CONCLUSIONS

It was compared the recommendations of the diallel analysis crosses with the process end the result of inbreeding generations. Recommended crosses by the diallel analysis were different from those obtained after generations of self-fertilization and selection. However, two of the four selected parents were matched. Considerations were made about the accuracy of diallel analysis and the need to validate in practice methodologies that theoretically are of great value.

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